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# Optimal investment for a retirement plan with deferred annuities allowing for inflation and labour income risk



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#### ABSTRACT

We construct an optimal investment portfolio model for an individual investor saving in a retirement plan. The investor earns stochastic labour income with both permanent and temporary shocks, and has access to equity, conventional bond, inflation-indexed bond and cash, as well as two types of deferred annuities: nominal and inflation-protected. The objective function consists of power utility in terms of real retirement income from the annuities as well as bequest from remaining wealth in tradable securities. Asset returns are represented by a vector autoregressive model underpinned by Nelson-Siegel real and nominal yield curves. The optimization problem is solved numerically using multi-stage stochastic programming with a hybrid scenario structure combining a scenario tree with scenario fans. Our numerical results show that deferred annuities are bought early and in increasing amounts during the working lifetime of the investor, with portfolio risk declining with age. Welfare is diminished by 40% if deferred annuities are not available. Inflation-protected deferred annuities are marginally more important in the presence of real labour income risk, but nominal deferred annuities are bought as a cheaper alternative if real yields are low or negative. Portfolio composition and annuity allocation vary depending on financial market expectations, but our central result about the importance of deferred annuities is robust to a variety of financial market conditions.

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#### 1. Introduction

A life annuity is a product which is sold by life insurers to individuals and which makes a regular stream of payments to the annuity-holder while alive. An annuity is termed an "immediate annuity" if it starts paying out from the time that it is purchased. The insurer can also design an annuity which has a deferment period between the time that the annuity is purchased and its first payout date, and this is known as a "deferred annuity" (or deferred income annuity). Annuities are a key product in retirement planning since retirees face longevity risk, i.e. the risk that they will outlive their savings. By purchasing an annuity, individuals transfer longevity risk to life insurers, which can pool this risk by writing a large portfolio of annuities. Unsurprisingly, the annuity market is very large. In 2017 in the U.S., sales of annuities totaled \$203.5 billion, of which deferred annuities accounted for \$2.2 billion (Chen, Haberman, & Thomas, 2019). In addition to the retail market, a type of deferred annuity is also provided by social security and defined benefit pension plans, in which case governments and employers, rather than insurers, back the retirement benefits.

All life annuities help individuals by removing longevity risk, but deferred annuities have a specific role in reducing longevity risk at very old ages for retirees (Ezra, 2016; Scott, 2008). Deferred annuities are also useful for younger working individuals as they can be used to secure retirement income well before retirement (Horneff, Maurer, & Rogalla, 2010; Maurer, Mitchell, Rogalla, & Kartashov, 2013). Annuity payments may be fixed in nominal terms or may be linked to a specific index, for example an index of inflation. Inflation-indexed (or inflation-protected) deferred annuities can be particularly useful since payments may continue for a long time in the future, and retirees' purchasing power should be protected (Merton, 2014).

Although deferred annuities are potentially very useful in retirement, there is little research on optimal retirement planning with these annuities. Target-date funds in the U.S. are authorised to include deferred annuities in 401(k) pension plans (U.S. Treasury Department, 2014), but in practice many pension plans implement simple so-called glide-path strategies (Donaldson, Kinniry, Maciulis, Patterson, & DiJoseph, 2015) which disregard deferred annuities. Glide paths reduce equity allocation and increase bond

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allocation as retirement draws nearer, but such a strategy does not maximize utility in terms of retirement income in real terms (Merton, 2014).

Only a few studies look at optimal investment with deferred annuities prior to retirement. Horneff et al. (2010) and Maurer et al. (2013) find that it is optimal to start purchasing deferred annuities from about age 40 with continual purchase of up to about 80% of wealth at retirement. They assume one risk-free asset with constant interest rate, one risky asset, and stochastic labour income. A similar conclusion is reached by Konicz and Mulvey (2013) in a different setting with full consumption of annuity income during retirement. Huang, Milevsky, and Young (2017) seek the conditions for purchase of deferred annuities when interest rates are mean-reverting but their analysis is limited because there is neither investment portfolio optimization nor labour income.

Unlike the sparse research on the use of deferred annuities in optimal investment prior to retirement, there is considerably more research on immediate annuities on and after retirement. Koijen, Nijman, and Werker (2011) calculate the optimal allocation to three types of immediate annuities (nominal, inflationprotected and variable) when wealth is fully annuitized at retirement. They also find the optimal pre-retirement consumption and investment to hedge the optimal annuity portfolio at retirement. Another strand of annuity research concerns the timing of purchase of immediate annuities. Horneff, Maurer, Mitchell, and Dus (2008) and Horneff, Maurer, Mitchell, and Stamos (2009) solve numerically for the optimal annuitization and investment decisions when nominal immediate annuities are available in retirement and variable immediate annuities, stocks and bonds are available before retirement. They find that annuity holdings increase over an individual's lifetime. They also obtain the typical life-cycle result that equity allocation falls while bond allocation increases over time.

An important aspect of retirement planning is labour income. Most occupational retirement plans require employees to make a pension contribution which is a fixed proportion of their income. Viceira (2001), Cocco, Gomes, and Maenhout (2005) and Benzoni, Collin-Dufresne, and Goldstein (2007) examine how risks to labour income influence optimal investment choices over an individual's lifetime. An individual, whose labour income stream is deterministic and non-tradable during her working lifetime, can be regarded as holding a risk-free coupon bond. Therefore, the investor's optimal allocation to a risky asset is higher than if she had an uncertain labour income, and it decreases over time. On the other hand, if labour income is volatile and positively correlated with the risky asset return, the fraction of wealth invested in the risky asset decreases (Cocco et al., 2005). A similar result is found by Benzoni et al. (2007) who show that cointegration between labour income and stock dividends can lead the investor to take a short position in the risky assets. However, if labour income risk is idiosyncratic, risky asset allocation can be larger than in the absence of labour income (Viceira, 2001).

Our main contributions in this article are five-fold. First, we employ a rich and realistic model of financial markets which means that our results can be implemented by pension and financial planners. Previous studies on deferred annuities in individual portfolio optimization, such as by Horneff et al. (2009, 2010), Maurer et al. (2013), and Konicz and Mulvey (2013), use only a constant risk-free rate and a geometric Brownian motion-driven risky asset. Koijen et al. (2011) employ a model with time-varying equity return and a full term structure, which they argue is critical to long-term portfolio allocation. In this paper, we use a vector autoregressive model underpinned by a Nelson–Siegel model of the term structure of nominal and real interest rates, similar to Konicz, Pisinger, and Weissensteiner (2015). However, we go further than Koijen et al. (2011) and Konicz et al. (2015) in that we include deferred annuities whereas they do not.

Second, we model a richer set of instruments and features, again making our model implementable by retirement planners. In particular, we have inflation-indexed bonds as well as inflation-indexed annuities, both immediate and deferred. We assume that pension savers have stochastic labour income, with an inflation component, as would be the case in the real world. It is especially important to include inflation given the long-term nature of retirement planning. To our knowledge, no other study on deferred annuitization incorporates all these features. Further, short-selling is not allowed in our model. Koijen et al. (2011) find that large amounts of short-selling occur in their optimal solution (with immediate annuities only), but this cannot be implemented as individual investors do not have access to short positions within their long-only retirement funds.

Third, we use multi-stage stochastic programming (MSP) to contend with the numerous state variables that our more realistic setting imposes. Numerical dynamic programming is used by Koijen et al. (2011) but they fail to incorporate realistic features and constraints such as deferred annuities and no short sales. MSP is also used by Consigli, Iaquinta, Moriggia, di Tria, and Musitelli (2012), Dempster and Medova (2011), Konicz and Mulvey (2013, 2015) and Konicz et al. (2015) in individual retirement planning, but we extend these works either with a richer financial market model or with a richer set of instruments or both. MSP requires the use of scenarios, and we generate several scenario trees that are arbitrage-free. Furthermore, we introduce a new scenario structure which combines a scenario tree with scenario fans in order to track inflation scenarios which are used for projecting annuity income in retirement. We also devise an optimization model to construct scenarios of labour income shocks which are independent of the financial market scenarios.

Our fourth contribution in this article is to evaluate explicitly the welfare gains when individuals have immediate and deferred annuities at their disposal, in both nominally-fixed and inflationindexed varieties. This is important because it enables individual investors as well as their financial advisers to observe in monetary terms the advantage of including these instruments in their pension planning.

The fifth and final contribution of our paper is that we carry out sensitivity analyses over different historical periods. Financial market conditions change, particularly after momentous events such as the 2008 financial crisis, and we generate new scenario trees and find the optimal investment over different periods. This should reinforce financial planners' confidence in the model, and provide them with a robust tool with which to advise investors saving for retirement.

#### 2. Investment problem for a retirement plan

Assume that there is an individual, aged  $\delta$  years old at time 0, who makes regular contributions to a personal retirement plan until his retirement at time *T*. He lives to a maximum age  $\omega$  (the last age in an actuarial life table), so he cannot live beyond time  $\tau = \omega - \delta$ . The amount of the contributions is a fixed proportion  $\phi$  of his uncertain nominal labour income  $L_t$  (at time *t*) during in the pre-retirement period [0, *T*).

In the retirement fund, the individual can hold equity, nominal bond, inflation-linked bond, and cash. In the remainder of the paper, these financial assets are denoted using subscripts *E*, *B*,  $\widetilde{B}$  and *C* respectively, with the set of financial assets being  $\mathcal{F} = \{E, B, \widetilde{B}, C\}$ .

Withdrawals from the retirement fund are allowed but only to buy deferred annuities (DAs) which will pay out, if he is alive at retirement time *T*, every year from time *T* until he dies. Two types of annuities can be purchased: nominal and inflation-protected annuities, denoted by the subscripts *A* and  $\widetilde{A}$  respectively. The set of annuity products is  $\mathcal{A} = \{A, \widetilde{A}\}$ . Both types of annuities are irreversible contracts, so the individual can buy them from insurers, but not sell them back to insurers or on a secondary market.

There is no payout from the annuities if the individual dies before retirement. However, any remaining wealth in his retirement fund is bequeathed to his heirs. If the individual survives till retirement, then all of his accumulated wealth in the fund is used to purchase nominal and inflation-protected immediate annuities which pay out from the retirement date until death. An immediate annuity is merely a deferred annuity with a zero deferment period, so in the following we do not distinguish between immediate and deferred annuities.

We assume that the investor exhibits constant relative risk aversion (CRRA) with a power utility function  $u(t, x) = e^{-\rho t} (x^{1-\gamma})/(1-\gamma)$ , at time t, with  $\gamma > 0$  being a risk aversion coefficient,  $0 \le \rho \le 1$  being a time preference coefficient, and x representing either annuity income or wealth bequest, both of the latter being in real terms.

If the individual survives to retirement at time *T*, then utility is gained by retirement income *in real terms* (i.e. net of inflation), following the policy prescription of Merton (2014). Real income in retirement is denoted by  $l_t$  for  $t \in [T, \tau)$  and it is the nominal income from annuities deflated back to time 0 at the realized logrates of price inflation { $\Pi_t$ ,  $t \in (0, \tau]$ }. We assume implicitly that all income from the retirement plan is used for consumption, so utility  $u(t, l_t)$  is gained for annuity income, in real terms, every year in retirement.

If the individual dies before time *T*, then his wealth in the retirement plan (minus the value of deferred annuities purchased) is bequeathed to his estate, and utility is gained by the amount, in real terms, that is bequeathed. Real wealth bequeathed is  $W_t$  for  $t \in (0, T]$ , and it is the nominal value of wealth deflated back to time 0. A bequest parameter  $\kappa$ , to be introduced shortly, captures the relative importance of the bequest to retirement income.

During the pre-retirement period [0, T), the investor can dynamically adjust his investment portfolio and purchase deferred annuities in order to maximize the expected utility of real annuity income in retirement and of real bequest before retirement. Let the total number of units of nominal deferred annuity purchased by time *t* be  $X_{A,t}$  and, similarly, the total number of units of inflation-protected deferred annuity be  $X_{\tilde{A},t}$ . (Recall that the subscript *A* stands for nominal annuity and  $\tilde{A}$  for inflation-protected annuity.)

One unit of nominal annuity guarantees a nominal income of £1 annually in retirement until death. One unit of inflation-protected annuity also guarantees a nominal income of £1 annually in retirement until death, but the number of units is increased in line with price inflation every year. Thus, income from the inflation-protected annuity is perfectly correlated with price inflation. Let the prices of one unit of nominal and inflation-protected annuities be  $S_{A,t}$  and  $S_{\overline{A},t}$  respectively at time  $t \in [0, T]$ . Then, at time t, the investor pays  $S_{A,t}(X_{A,t} - X_{A,t-1})$  to buy the nominal annuity and  $S_{\overline{A},t}(X_{\overline{A},t} - X_{\overline{A},t-1}e^{\Pi_t})$  to buy the inflation-protected annuity, where  $\Pi_t$  is the log-rate of price inflation in year (t - 1, t). (In both cases,  $X_{A,-1} = X_{\overline{A},-1} \equiv 0$ .)

The investor can also buy and sell units of a cash fund, equity fund, nominal bond fund and inflation-linked bond fund. (Recall that these are denoted by *C*, *E*, *B* and  $\tilde{B}$  respectively.) Let  $X_{E,t}$  be the number of units of the equity fund held in the retirement plan at time *t*, and  $S_{E,t}$  be the price of equity units at time *t*. A corresponding notation holds for the cash, nominal bond and inflation-linked bond funds. Note that we assume perfect divisibility of assets and that fractional units can be held.

At time  $t \in [0, T)$ , the individual can adjust asset allocations of the retirement plan by deciding how much to hold in cash, equity, and bonds, and how many annuity units to buy. At the retirement date T, all financial wealth is sold and the investor de-

cides how many units of nominal and inflation-protected annuities to buy. The decision variable for the individual at time  $t \in [0, T]$  is therefore  $X_t = [X_{C,t}, X_{B,t}, X_{E,t}, X_{\tilde{B},t}, X_{A,t}, X_{\tilde{A},t}]'$ .

#### 3. Formulation for multi-stage stochastic programming

#### 3.1. Preliminary notation and definitions

Stochastic programming is a mathematical framework for optimization problems with uncertain scenarios. The scenarios can be economic, financial, and demographic. Both the state space and time are discretized in multi-stage stochastic programming (MSP). The multiple discrete-time points are known as stages. An MSP model is constructed in a nodal form by using state variables generated in a scenario tree. The scenario tree starts at the initial stage from a unique root node and it ends at the terminal stage with multiple leaf nodes. The root node branches out to a number of children nodes at the second time stage. Each child node itself branches out to further nodes at the third time stage, and so on, until the leaf nodes are reached. Every node, except the root node, has a unique parent node. A scenario is the connected path through a series of parent nodes from a leaf node to the root node. In general, the scenario tree is non-recombining.

We set out here some notation pertaining to the scenario tree. For convenience, this notation is summarized in Table 1. The scenario tree is depicted in schematic form in Fig. 1. The root node is located at the first stage and is denoted by  $n_0$ . The set of all nodes in the scenario structure is  $\mathcal{N}$ , and  $\mathcal{N}_t$  is the set of nodes at time *t*. Thus,  $\mathcal{N}_0 = \{n_0\}$  contains the root node only,  $\mathcal{N}_\tau$  is the set of leaf nodes, and  $\mathcal{N} = \bigcup_{t \in [0,\tau]} \mathcal{N}_t$ . A node  $n \neq n_0$  branches off a parent node, denoted by  $n^-$ , which may itself have its own parent node  $n^{--}$ , etc. It is convenient to denote by  $s_n$  the set of all predecessor nodes of node n, i.e.  $s_n = \{n, n^-, n^{--}, \dots, n_0\}$ . A node  $n \notin \mathcal{N}_{\tau}$  forks into a set of children nodes, denoted by  $\{n^+\}$  at the next stage, and these children nodes may themselves have their own children nodes  $\{n^{++}\}$  at the following stage, etc. The time between each stage may vary, in general, but it is fixed in our model and is denoted by  $\Delta t$ . The unconditional probability that a node *n* occurs is  $\mathbf{pr}_n$  and the conditional probability that a node *n* occurs given its parent node  $n^-$  is  $pr_n$ .

We make use of a hybrid structure of scenario tree followed by scenario fans, as illustrated in Fig. 1. The scenario tree consists of multiple branches from every node and it spans the decision period [0, T], i.e. the period up to retirement during which investment and annuitization decisions are made. The scenario fans consist of only two branches from every node at retirement, and they span the projection period  $(T, \tau]$ , i.e. the post-retirement period. It is important to highlight that decisions are only made up to retirement, and the relevant scenarios are captured in the scenario tree. The scenario fans after retirement are only used to project inflation forward. They are required because income is received during retirement but it is the utility of retirement income in real terms (i.e. net of inflation) that must be evaluated. As explained by Dupačová, Consigli, and Wallace (2000), the scenario fan structure can be used in the special case in which the probability distributions of supporting variables are only affected by decisions made before the scenarios.<sup>1</sup>

In the following, survival and death probabilities are represented using standard actuarial notation. The probability that a  $\delta$ -year old person survives until age  $\delta + t$  is denoted by  $_t p_{\delta}$ . The probability that a ( $\delta + t$ )-year old person dies over the following

<sup>&</sup>lt;sup>1</sup> The combined structure of the scenario tree and scenario fans, depicted in Fig. 1, still satisfies the non-anticipative condition. Here, the supporting variable is the inflation process to project annuity income, and decisions are all made before starting the scenario fan.

#### Table 1

/ariables and param	eters for the	MSP optimization	problem.

<u>Sets</u>	
$\mathcal{N} = \bigcup_{t \in [0,\tau]} \mathcal{N}_t$	All nodes in the hybrid scenario structure
$\mathcal{N}_t$	All nodes at time t
$\mathcal{A} = \{A, \widetilde{A}\}$	Set of annuities guaranteeing fixed-nominal (A) and inflation-protected ( $\widetilde{A}$ ) payments
$\mathcal{F} = \{C, B, \widetilde{B}, E\}$	Set of financial assets: cash (C), conventional bond (B), inflation-linked bond $(\tilde{B})$ , equity (E)
$s_n = \{n, n^-, n^{}, \dots, n_0\}$	All predecessor nodes of node $n$ , including itself, up to root node $n_0$
$\{n, n^+, n^{++}, \dots\}$	All successor nodes of node n, including itself
Decision variables	
$X_{in}^{buy}, i \in \mathcal{A} \cup \mathcal{F}$	The unit number of asset <i>i</i> to buy at node <i>n</i>
$X_{i,n}^{sell}, i \in \mathcal{F}$	The unit number of asset <i>i</i> to sell at node <i>n</i>
Xn	Vector collecting all the buy and sell decision variables at node <i>n</i>
$X_{i,n}, i \in \mathcal{A} \cup \mathcal{F}$	The unit number of asset $i$ held at node $n$ after rebalancing
Other main variables	
In	Real retirement income at node <i>n</i>
W <sub>n</sub>	Real financial wealth at node $n$ (excludes the value of annuities)
<u>Parameters</u>	
$_t p_{\delta}$	Probability that a $\delta$ -year old person survives until age $\delta + t$
$\Delta t q_{\delta+t}$	Probability that a $(\delta + t)$ -year old person dies over the following $\Delta t$ years
<b>pr</b> <sub>n</sub>	Unconditional probability that a node <i>n</i> occurs
$S_{i,n}, i \in \mathcal{A} \cup \mathcal{F}$	Nominal price of asset <i>i</i> at node <i>n</i>
L <sub>n</sub>	Nominal labour income per annum at node <i>n</i>
$\Pi_n$	Inflation log-rate over a $\Delta t$ -long time interval ending at node $n$
$\Delta t, T, \tau$	Portfolio holding period, retirement date, maximum time respectively (all in years)
δ	Investor's starting age at $t = 0$
$(\gamma, \rho, \kappa)$	Investor's preference parameters; risk aversion, time preference, and bequest motive respectively
<i>w</i> <sub>0</sub>	Current nominal wealth in cash at the root node $n_0$ before contribution and rebalancing
$\phi$	Fixed contribution rate, as a proportion of nominal labour income $L_n$ , to the retirement plan



**Fig. 1.** A hybrid structure of a scenario tree (from time 0 to *T*) and scenario fans (from T + 1 to  $\tau$ ).

 $\Delta t$  years is denoted by  $_{\Delta t}q_{\delta+t}$  (typically abbreviated to  $q_{\delta+t}$  when  $\Delta t = 1$ ).

#### 3.2. Optimization problem

The optimization problem for the investor in a retirement plan can be formulated on the scenario structure as a multi-stage stochastic programming (MSP) problem as follows. Let  $X_{i,n}$  be the number of units of asset  $i \in A \cup F$  held at node n in the scenario structure. We separate buy and sell decisions:  $X_{i,n}^{buy}$  is the number of units of asset *i* to buy at node *n*, and  $X_{i,n}^{sell}$  is the number of units of asset *i* to sell at node *n*. Annuities cannot be sold, so the decision variable at node *n* is  $X_n = \begin{bmatrix} X_{C,n}^{buy}, X_{B,n}^{sell}, X_{B,n}^{buy}, X_{E,n}^{sell}, X_{\widetilde{B},n}^{buy}, X_{\widetilde{B},n}^{sell}, X_{\widetilde{B},n}^{$ 

The objective function, budget constraints, and variable constraints for the retirement planning problem are given in nodal form suitable for MSP by the equations below (the equations in the usual time representation are reproduced in the online supplementary appendix, section S-1):

$$\max_{\{X_n, n \in \{\mathcal{N}_t, t \in [0,T]\}\}} \left[ \sum_{t \in [T,\tau)} \sum_{n \in \mathcal{N}_t} {}_t p_{\delta} u(t, I_n) \mathbf{pr}_n + \sum_{t \in [0,T)} \sum_{n \in \mathcal{N}_{t+\Delta t}} {}_t p_{\delta \Delta t} q_{\delta+t} \kappa^{\gamma} u(t + \Delta t, W_n) \mathbf{pr}_n \right],$$
(1a)

s.t. 
$$I_n = \left(\sum_{i \in \mathcal{A}} X_{i,n}\right) \exp\left(-\sum_{m \in S_n} \Pi_m\right)$$
 for  $n \in \{\mathcal{N}_t, t \in [T, \tau)\},$ 
(1b)

$$W_n = \sum_{i \in \mathcal{F}} X_{i,n^-} \cdot S_{i,n} \exp\left(-\sum_{m \in S_n} \Pi_m\right) \text{ for } n \in \{\mathcal{N}_t, t \in (0,T]\},$$
(1c)

$$\mathbb{1}_{\{n=n_0\}} w_0 + \mathbb{1}_{\{n \notin \mathcal{N}_T\}} \phi \ L_n \ \Delta t + \sum_{i \in \mathcal{F}} X_{i,n}^{sell} \ S_{i,n}$$
$$= \sum_{i \in \mathcal{A} \cup \mathcal{F}} X_{i,n}^{buy} \ S_{i,n} \quad \text{for } n \in \{\mathcal{N}_t, t \in [0, T]\},$$
(1d)

$$X_{i,n} = \mathbb{1}_{\{n \neq n_0\}} X_{i,n^-} + X_{i,n}^{buy} - X_{i,n}^{sell}$$
  
for  $i \in \mathcal{F}$  and  $n \in \{\mathcal{N}_t, t \in [0, T]\},$  (1e)

$$X_{A,n} = \mathbb{1}_{\{n \neq n_0\}} X_{A,n^-} + \mathbb{1}_{\{n \in \{\mathcal{N}_t, t \in [0,T]\}\}} X_{A,n}^{buy}$$
  
for  $n \in \{\mathcal{N}_t, t \in [0, \tau]\},$  (1f)

$$X_{\tilde{A},n} = \mathbb{1}_{\{n \neq n_0\}} X_{\tilde{A},n^-} e^{\Pi_n} + \mathbb{1}_{\{n \in \{\mathcal{N}_t, t \in [0,T]\}\}} X_{\tilde{A},n}^{buy}$$
  
for  $n \in \{\mathcal{N}_t, t \in [0, \tau]\},$  (1g)

$$X_{i,n} = X_{i,n}^{\text{buy}} = 0 \quad \text{for } i \in \mathcal{F} \text{ and } n \in \mathcal{N}_T,$$
(1h)

$$X_{i,n}, X_{i,n}^{buy}, X_{i,n}^{sell} \ge 0 \quad \text{for } i \in \mathcal{F} \text{ and } n \in \{\mathcal{N}_t, t \in [0, T]\},$$
(1i)

$$X_{i,n}^{buy} \ge 0 \quad \text{for } i \in \mathcal{A} \text{ and } n \in \{\mathcal{N}_t, t \in [0, T]\}.$$
(1j)

In Eq. (1a) above, the decision variables over which the expected utility is maximized are the portfolio and annuity purchase decisions over the planning horizon [0, T]. It is implicit in the objective function in Eq. (1a) that summations occur over the time stages in the scenario tree during the investor's lifetime  $t \in [0, \tau)$ . The decision variable must be chosen at every node n in the scenario tree component of the scenario structure,  $\{N_t, t \in [0, T]\}$ . Eq. (1b) shows the real retirement income  $I_n$  at time  $t \in [T, \tau)$  during retirement, this being the nominal income from annuities deflated back to time 0. Recall that the number  $X_{\tilde{A},n}$  of inflation-protected annuity units increases in line with price inflation, whereas the corresponding number for nominal units does not. Wealth  $W_n$  in Eq. (1c) is evaluated in real terms and does not include the value of annuities, so it is the inflation-adjusted total value of financial assets before rebalancing the portfolio at node n.

The cash balance constraint in Eq. (1d) sets off cash inflows against outflows. At the root node, wealth is initialized at the non-random amount  $w_0$  specified on the l.h.s. of Eq. (1d). The incidence of cash flows in our model is such that contributions occur in advance every  $\Delta t$  years. A financial asset inventory constraint appears in Eq. (1e) and tracks the number of units of cash, equity, nominal and inflation-indexed bond funds held at node *n*. The

annuity inventory constraints in Eqs. (1f) and (1g) allow the investor to buy the two types of annuities during the pre-retirement planning period as well as at retirement, and they track the number of units of annuities for the whole lifetime  $[0, \tau)$ . Notice that the number of units of inflation-protected annuities in Eq. (1g) increases in line with inflation every year. At retirement time *T*, all assets are sold to purchase nominal and inflation-protected immediate annuities, and constraint Eq. (1h) guarantees this full annuitization. Eq. (1i) represents the no short-sales constraint, while Eq. (1j) ensures that annuities can only be bought and not sold.

#### 3.3. Scenario generation for financial and annuity markets

There are two related features of the scenario tree which call for a modelling decision: the number of stages in the tree and the branching factor at every stage. (The branching factor is the number of children nodes that each node has.) A "curse of dimensionality" occurs in MSP as the number of stages increases and the scenario tree becomes 'denser' (Dupačová et al., 2000; Shapiro, Dentcheva, & Ruszczynski, 2009). This corresponds to a similar curse in numerical dynamic programming when the number of state variables increases. In MSP, there is a computational limit to the number of stages and branching factors in the scenario tree.

A common strategy is to use stages of increasing lengths with decreasing branching factors along the tree, as used for example by Consigli et al. (2012). We do not employ this strategy here for two reasons. First, branching factors cannot decrease beyond the minimum number of branches required to avoid arbitrage (Geyer, Hanke, & Weissensteiner, 2010) and to match the required moments of the conditional distributions of the variables stored at every node in our optimization problem (Høyland & Wallace, 2001). Second, we use a regular time interval between stages to replicate the regular financial reviews that an individual may have with a financial planner during their financial lifecycle, and the regular portfolio rebalancing that then occurs. The optimal investment solution that we obtain is then an approximation to optimal intertemporal investment without the distortion that would be caused by an uneven time-interval effect.<sup>2</sup>

To generate scenario trees for a portfolio optimization problem, scenario reduction and state aggregation are not suitable methods (Geyer et al., 2010). Scenario reduction methods do not admit no-arbitrage conditions explicitly, while state aggregation involves only the risk-neutral measure, not the real-world measure which should be used in a portfolio optimization problem. To generate the scenario tree, we combine the sequential approach of Høyland and Wallace (2001) with the moment matching method (Klaassen, 2002).

Details of the scenario generation procedure are given in Section S-2 of the online supplementary appendix. In brief, we have 10 state variables at every node, we use a regular stage interval of  $\Delta t = 5$  years and we have 6 stages (5 periods), so that the scenario tree spans 25 years in the decision period up to retirement (this is the period [0, T] with T = 25 in Fig. 1). Given the values of the variables on any particular node n, the moments of the conditional distribution of these state variables after  $\Delta t = 5$  years can be calculated using a suitable financial model fitted to market data. (The model will be described in Section 4.) The values of the state variables on the children nodes  $\{n^+\}$  are then determined

<sup>&</sup>lt;sup>2</sup> We thank a reviewer for suggesting that stages of increasing lengths, with a short period of 1 year between the first two stages, could be useful in a practical setting. For example, the model could be used in a receding horizon control fashion where it is fully solved but only the first step is implemented. It is then solved again the following year, when the individual is one year older, and again only the first step is implemented, etc. until retirement.

by matching the first 4 moments of the conditional distribution. There are 85 moment specifications  $(10 \times 4 \text{ central moments and } 10(10-1)/2 = 45 \text{ covariances})$ . Based on this, Høyland and Wallace (2001) suggest that there should be a minimum of 8 branching factors. Starting from the root node, we can evaluate the state variables at each node sequentially throughout the scenario tree. This procedure gives a close fit to market data as is demonstrated in Section S-2 of the online supplementary appendix.

Another consideration relevant to scenario generation is that the scenarios are arbitrage-free. The following procedure is used to preclude arbitrage:<sup>3</sup>

- Step 1. Given a node  $n \in \{N_t, t \in [0, T)\}$ , calibrate the first four conditional moments of the subtree branching from this node.
- Step 2. Check if the generated scenarios preclude arbitrage opportunities among cash, nominal bond, inflation-linked bond, and equity funds: see Klaassen (2002). Return to Step 1 if any arbitrage opportunity is found.
- Step 3. Check if each of the generated scenarios has conditional moments over the next stage that are within no-arbitrage bounds: see Geyer, Hanke, and Weissensteiner (2014).
- Step 4. Repeat Step 1 to Step 3, if Step 3 meets the alwaysarbitrage bound.

Step 2 can be subsumed within Step 1. We proceed in a sequential way and apply the above procedure from the first to the penultimate stage. To identify arbitrage opportunities among the four financial assets (cash, nominal bond, inflation-linked bond, and equity) in Steps 2 and 3, we use the two methods of Klaassen (2002) for two arbitrage types *ex-post* and the method of Geyer et al. (2014) for no-arbitrage bounds *ex-ante*.

# 3.4. Scenarios for pre-retirement labour income and post-retirement inflation

In the pre-retirement period (this is the period [0, T) in Fig. 1), labour income is earned by the individual. In order to incorporate the effects of stochastic labour income into the retirement planning problem, we generate labour income scenarios superposed upon the financial market scenarios described in the previous section. We can generate real labour income scenarios which are independent of the asset return scenarios during the individual investor's working lifetime [0, T) because there are enough degrees of freedom in the eight branches of each node within the financial market scenarios, when the ten state variables are reduced to four financial assets returns for cash, equity, nominal bond and inflation-linked bond funds. Visiting each node in the working period of the scenario structure, we generate permanent real labour income shocks and temporary real labour income shocks on the eight children nodes. Details of the labour income model are given in Section 4.4 below and the relevant scenario generation is discussed in Section S-3 of the online supplementary appendix.

In the post-retirement period (this is the period  $[T, \tau]$  in Fig. 1), there is no labour income, of course, but annuity income is earned. This income must be evaluated *in real terms*, i.e. deflated back to the root node  $n_0$  at time 0. We use scenario fans to project inflation forward during retirement. We emphasize again that no decision is made in the post-retirement period. Log-rates of inflation are Normally distributed in our model (see Section 4.3 below), consequently only two moments need to be matched to create scenario fans. For the scenario fans, we have an initial branching factor of 2 and we choose two points at one conditional standard

deviation on either side of conditional mean inflation on each node in  $N_T$ , i.e. at retirement (Dupačová et al., 2000).

In the scenario tree part of the hybrid scenario structure that we use, there are eight children nodes for every node over the first six stages (five periods), giving  $8^5 = 32$ , 768 scenarios in the scenario tree. This is then doubled to 65,536 in the overall hybrid scenario structure because of the scenario fans. Since post-retirement inflation is required only for the purpose of deflating retirement income back to the root node  $n_0$  at time 0, 65,536 scenarios for inflation is an ample number to capture Normally distributed logrates of inflation, conditioned on the root node. Finally, intervals of one year are used along the fans, from time T + 1 = 26 to the terminal date  $\tau = 80$  in Fig. 1. (As we indicate later, the terminal date will correspond to a maximum age in a mortality table.)

#### 4. Financial assumptions

#### 4.1. Term structure of interest rates

The Nelson–Siegel model is chosen to model the real and nominal yield curves along with a vector autoregressive (VAR) model for stochastic cash, bond, equity returns, and inflation rates. The Nelson–Siegel model is parsimonious and known to avoid the overfitting problem and to return better out-of-samples predictions than affine term structure models (Diebold & Li, 2006). Ferstl and Weissensteiner (2011) combine the Nelson–Siegel formulation proposed by Boender, Dert, Heemskerk, and Hoek (2008) with the VAR model. Our model therefore incorporates asset return predictabilities and produces a seamless yield curve for pricing not only the cash and nominal and inflation-indexed bond funds, but also nominal and inflation-protected annuities.

The entire nominal yield curve is determined by a fitted Nelson–Siegel model with three time-varying parameters<sup>4</sup>,  $\beta_t = [\beta_{1,t}, \beta_{2,t}, \beta_{3,t}]'$ . For the real yield curve, we use the notation  $\tilde{\beta}_t = [\tilde{\beta}_{1,t}, \tilde{\beta}_{2,t}, \tilde{\beta}_{3,t}]'$ . The Nelson–Siegel model for the *s*-year nominal spot rate at time *t* is as follows:

$$y(\boldsymbol{\beta}_{t},s,\lambda) = \beta_{1,t} + (\beta_{2,t} + \beta_{3,t}) \left(\frac{1 - e^{-\lambda s}}{\lambda s}\right) - \beta_{3,t} e^{-\lambda s}, \qquad (2)$$

where the scaling parameter  $\lambda$  is a constant. A corresponding equation holds for the real spot rate.

#### 4.2. Time-varying investment opportunities

In order to incorporate predictabilities of asset returns and a pair of three Nelson–Siegel parameters ( $\beta_t$  and  $\tilde{\beta}_t$ ), we use a VAR(1) model (for details, see Barberis, 2000; Campbell, Chan, & Viceira, 2003). Specifically, we use the combined approach of Konicz et al. (2015) to model interest rates, equity returns and inflation rates. Here,  $r_t$  is monthly log-returns on the equity fund. Monthly inflation log-rates are denoted by  $\pi_t$ . Our VAR model is given by

$$z_t = \Phi_0 + \Phi_1 z_{t-1} + v_t, \tag{3}$$

where  $z_t = [r_t, \pi_t, \beta_{1,t}, \beta_{2,t}, \beta_{3,t}, \tilde{\beta}_{1,t}, \tilde{\beta}_{2,t}, \tilde{\beta}_{3,t}]'$ . The accumulated return  $R_{E,t}$  on the equity fund, defined near Eq. (6) as the log-return from time  $t - \Delta t$  to t, is simply a sum of the monthly log-returns. The accumulated inflation  $\Pi_t$ , defined near Eq. (8) and Eq. (12), is a sum of the monthly inflation rates. In Eq. (3), the intercept term  $\Phi_0$  is a column vector. The slope term  $\Phi_1$  is a  $8 \times 8$ 

<sup>&</sup>lt;sup>3</sup> We are grateful to a reviewer for suggesting the alternative method set out by Consiglio, Carollo, and Zenios (2016).

<sup>&</sup>lt;sup>4</sup> In the Nelson–Siegel model, the long interest rate is given by  $\lim_{s\to\infty} y(\beta_t, s) = \beta_{1,t}$  and the short interest rate is  $\lim_{s\to0} y(\beta_t, s) = \beta_{1,t} + \beta_{2,t}$ . The parameters  $\beta_{1,t}$ ,  $\beta_{2,t}$ , and  $\beta_{3,t}$  determine level, slope, and curvature of the yield curve respectively (Boender et al., 2008).

Table 2							
Estimated	parameters	and	t-statistics	for	the	VAR(1)	model.

	$\Phi_0$	$\Phi_1$								
		$r_{t-1}$	$\pi_{t-1}$	$\beta_{1,t-1}$	$\beta_{2,t-1}$	$\beta_{3,t-1}$	$ ilde{eta}_{1,t-1}$	$ ilde{eta}_{2,t-1}$	$ ilde{eta}_{3,t-1}$	
r <sub>t</sub>	-0.0034	-0.0013	-0.0584	0.1265	-0.0425	-0.0342	0.1064	-0.0831	-0.0658	
	(-0.3415)	(-0.0251)	(-0.1051)	(0.4051)	(-0.2886)	(-0.3130)	(0.2384)	(-0.4921)	(-0.4000)	
$\pi_t$	0.0021	0.0016	0.0474	0.0288	0.0727	0.0026	-0.0690	-0.0844	-0.0157	
	(2.2691)	(0.3493)	(0.9413)	(1.0177)	(5.4389)	(0.2652)	(-1.7055)	(-5.5114)	(-1.0493)	
$\beta_{1,t}$	0.0031	-0.0021	-0.0402	0.9246	0.0301	0.0046	0.0280	-0.0155	0.0242	
	(2.9330)	(-0.3987)	(-0.6907)	(28.2607)	(1.9470)	(0.3996)	(0.5994)	(-0.8773)	(1.4043)	
$\beta_{2,t}$	0.0047	0.0156	0.0264	-0.1462	0.9498	-0.0243	0.1527	-0.0422	-0.0367	
	(3.2698)	(2.1473)	(0.3332)	(-3.2800)	(45.1243)	(-1.5571)	(2.3975)	(-1.7493)	(-1.5619)	
$\beta_{3,t}$	0.0043	-0.0057	-0.0123	-0.1073	-0.0930	0.8791	0.2206	0.0446	0.0040	
	(1.3922)	(-0.3662)	(-0.0719)	(-1.1206)	(-2.0568)	(26.2582)	(1.6133)	(0.8625)	(0.0800)	
$\tilde{\beta}_{1,t}$	0.0007	0.0003	-0.0105	-0.0220	0.0028	-0.0047	1.0178	0.0014	0.0169	
	(1.2741)	(0.1043)	(-0.3342)	(-1.2418)	(0.3383)	(-0.7583)	(40.2434)	(0.1499)	(1.8160)	
$\tilde{\beta}_{2,t}$	0.0028	-0.0141	0.3517	-0.1208	-0.0159	-0.0385	0.1352	0.9486	0.0132	
	(2.2349)	(-2.2155)	(5.0743)	(-3.1007)	(-0.8665)	(-2.8284)	(2.4294)	(45.0456)	(0.6422)	
$\tilde{\beta}_{3t}$	-0.0002	0.0082	-0.2255	0.0832	0.0234	0.0500	-0.1611	-0.0154	0.8587	
1 5,1	(-0.1062)	(0.8469)	(-2.1284)	(1.3969)	(0.8337)	(2.4032)	(-1.8940)	(-0.4785)	(27.3676)	

Monthly data of FTSE 100 from Bloomberg, retail price index from Office for National Statistics and fitted yield curves from Bank of England respectively are used from January 1985 to June 2017 ( $\lambda = 0.1519$  and  $\tilde{\lambda} = 0.2508$  for the Nelson–Siegel nominal and real yield curves model); t-statistics in parentheses.  $R^2$ : 0.0114 ( $r_t$ ), 0.1396 ( $\pi_t$ ), 0.9484 ( $\beta_{1,t}$ ), 0.9665 ( $\beta_{2,t}$ ), 0.8431 ( $\beta_{3,t}$ ), 0.9881 ( $\tilde{\beta}_{1,t}$ ), 0.92626 ( $\tilde{\beta}_{2,t}$ ), 0.8373 ( $\tilde{\beta}_{3,t}$ ).

Table 3

Standard deviations and cross correlations of residuals of the VAR(1) model.

	r <sub>t</sub>	$\pi_t$	$\beta_{1,t}$	$\beta_{2,t}$	$\beta_{3,t}$	$ ilde{eta}_{1,t}$	$ ilde{eta}_{2,t}$	$ ilde{eta}_{3,t}$
$ \sqrt{diag(\Sigma_z)} $ $ r_t $ $ \pi_t $ $ \beta_{1,t} $ $ \beta_{2,t} $ $ \beta_{3,t} $ $ \beta_{1,t} $ $ \beta_{2,t} $ $ \beta_{2,t} $	0.0442 1.0000	0.0040 -0.0192 1.0000	0.0046 0.1104 0.0161 1.0000	0.0063 -0.0786 0.0191 -0.3770 1.0000	0.0135 -0.1136 0.0665 -0.3500 0.5465 1.0000	0.0025 -0.0587 0.0545 0.3960 0.2576 0.3653 1.0000	0.0055 -0.0387 -0.0086 -0.0562 0.2980 0.0040 -0.0936 1.0000	0.0084 -0.0261 -0.0378 -0.1534 0.0509 0.2172 -0.4195 -0.3501
$p_{3,t}$								1.0000

Table 4

Unconditional expected mean  $\mu_{zz}$  and standard deviation  $\sigma_{zz}$  of the VAR(1) model.

	r <sub>t</sub>	$\pi_t$	$\beta_{1,t}$	$\beta_{2,t}$	$\beta_{3,t}$	$ ilde{eta}_{1,t}$	$ ilde{eta}_{2,t}$	$ ilde{eta}_{3,t}$
$\mu_{zz} \sigma_{zz}$	0.0022	0.0027	0.0392	-0.0283	0.0135	-0.0012	-0.0191	0.0210
	0.0445	0.0043	0.0219	0.0396	0.0341	0.0293	0.0206	0.0231

coefficient matrix of the VAR model. The error term  $v_t$  is a column vector of i.i.d. innovations ~ N(0,  $\Sigma_z$ ), where  $\Sigma_z = \mathbb{E}[vv']$ .

Given fixed values of  $\lambda$  and  $\tilde{\lambda}$ , if all eigenvalues of  $\Phi_1$  have moduli less than one, the stochastic process in Eq. (3) is stable with the unconditional expected mean  $\mu_{zz}$  and covariance  $\Gamma_{zz}$  of  $z_t$  in the steady state:

$$\mu_{zz} = (I - \Phi_1)^{-1} c \tag{4}$$

$$\operatorname{vec}(\Sigma_{zz}) = (I - \Phi_1 \otimes \Phi_1)^{-1} \operatorname{vec}(\Sigma_z), \tag{5}$$

where *I* is an identity matrix and the operator  $\otimes$  is the Kronecker product and *vec* is a vectorisation function, which transforms the  $8 \times 8$  matrix  $\Sigma_z$  into a  $8^2 \times 1$  vector.

Using historical nominal and real yield curves from the Bank of England from January 1985 to June 2017 with 0.5 to 25-year spot rates, monthly FTSE 100 data from Bloomberg, and retail price index as an inflation rate measure from the Office for National Statistics over the same period, the VAR model is stable. We choose the Nelson–Siegel scale parameters of  $\lambda = 0.1519$  and  $\tilde{\lambda} = 0.2508$  such that the sum of squared errors between the fitted Nelson–Siegel and historical yields is minimized. Our estimates for  $\Phi_0$  and  $\Phi_1$  in Eq. (3), along with *t*-statistics, are collected in Table 2. The level

of  $R^2$  for the equity return component is low, so it is difficult to confirm the existence of return predictability in the U.K. equity market. Table 3 exhibits the standard deviations and correlations of the residuals. Table 4 presents the unconditional expected mean  $\mu_{zz}$  and standard deviation  $\sigma_{zz} = \sqrt{diag(\Sigma_{zz})}$  of  $z_t$  at the steady state. The left hand panel of Fig. 2 displays the nominal term structure at the steady state, and it is clearly upward sloping and concave (as is the steady-state *real* term structure, not shown here). The right hand panel of Fig. 2 shows eight different stochastic realizations of the nominal term structure after 5 years.

#### 4.3. Price dynamics of bonds and annuities

The investor rebalances his portfolio and buys deferred annuities at regular intervals of length  $\Delta t$  years during his retirement planning period [0, T]. There are  $N \in \mathbb{N}$  such regular intervals, i.e.  $T = N\Delta t$ . Defining  $R_{i,t}$  as the accumulated log-return of financial asset  $i \in \mathcal{F}$  from time  $t - \Delta t$  to t, the price  $S_{i,t}$  of financial asset iis given by:

$$S_{i,t} = S_{i,t-\Delta t} \cdot \exp(R_{i,t}) \quad \text{for } i \in \mathcal{F},$$
 (6)

where  $S_{i,0} = 1$  without loss of generality.



Fig. 2. Left panel (a) shows the initial nominal term structure starting from the steady state. Right panel (b) shows different realizations of the nominal term structure after 5 years (the thicker the line, the likelier the occurrence).

The gross return of the nominal bond fund with a maturity of M years over a holding period of length  $\Delta t$  from time  $t - \Delta t$  to t is approximated by

$$R_{B,t} = M \cdot y(\boldsymbol{\beta}_{t-\Delta t}, M, \lambda) - (M - \Delta t) \cdot y(\boldsymbol{\beta}_t, M - \Delta t, \lambda).$$
(7)

The function  $y(\boldsymbol{\beta}_t, M, \lambda)$  denotes the *M*-year nominal spot rate at time *t*, determined by the Nelson–Siegel term structure model with a vector  $\boldsymbol{\beta}_t$  of parameters and a scale parameter  $\lambda$  to be specified shortly.

In a similar fashion, the gross return of the inflation-linked bond fund with a maturity of *M* years over a holding period of length  $\Delta t$  from time  $t - \Delta t$  to *t* is approximated by

$$R_{\tilde{B},t} = M \cdot y(\widetilde{\beta}_{t-\Delta t}, M, \widetilde{\lambda}) - (M - \Delta t) \cdot y(\widetilde{\beta}_{t}, M - \Delta t, \widetilde{\lambda}) + \Pi_{t},$$
(8)

where  $\Pi_t$  denotes the gross rate of inflation between  $t - \Delta t$  and t. The *M*-year real spot rate at time t is also denoted by the function y, but with parameters  $\tilde{\beta}_t$  and  $\tilde{\lambda}$ .

The gross interest rate on the cash fund from time  $t - \Delta t$  to t is defined simply by changing bond maturity M in Eq. (7) to  $\Delta t$ . The cash fund return is given by

$$R_{C,t} = \Delta t \cdot y(\boldsymbol{\beta}_{t-\Delta t}, \Delta t, \lambda).$$
(9)

Naturally, the cash return at time *t* does not depend on the current nominal spot rate  $y(\boldsymbol{\beta}_t, \Delta t, \lambda)$  at time *t*, but on the past spot rate  $y(\boldsymbol{\beta}_{t-\Delta t}, \Delta t, \lambda)$ . The price dynamics of the cash and bond funds are obtained by substituting  $R_{i,t}$  from Eq. (7), Eq. (8) and Eq. (9) into Eq. (6).

The fair actuarial price of a nominal deferred annuity contract which pays  $\pounds 1$  in nominal terms and in every year of retirement is

$$S_{A,t} = \sum_{s=T-t}^{\tau-t} {}^{s} p_{\delta+t} \cdot \exp\left(-s \cdot y(\boldsymbol{\beta}_{t}, s, \lambda)\right).$$
(10)

If t = T in Eq. (10) above, then the annuity is of course an immediate annuity. Eq. (10) above holds verbatim for an inflationprotected deferred annuity, except that *A* is replaced by  $\widetilde{A}$  and the yield is  $y(\widetilde{\beta}_t, s, \widetilde{\lambda})$ . We assume static pricing mortality rates here, and we also ignore loading factors (expenses).

#### 4.4. Labour income

We model exogenous labour income in accordance with Viceira (2001), Cocco et al. (2005) and Blake, Cairns, and Dowd (2007). An individual investor cannot change the amount of labour supply during his working lifetime [0, *T*). Real labour income  $\tilde{L}_t$  at time *t* is determined by three components: a deterministic growth function *f* of age  $\delta + t$ , a permanent shock  $\upsilon_t$  and a temporary shock

 $\varepsilon_t$ . The deterministic function f is a polynomial, which can match the hump shape of the age profile of real wages. The permanent labour income shock  $\upsilon_t$  can be either independent of, or correlated with, equity returns. Nominal labour income  $L_t$  at time t is the real labour income  $\tilde{L}_t$  inflated by two components: price inflation  $\Pi_t$  and real wage inflation  $G_t$ .

The real and nominal labour income processes of the investor are therefore given, respectively, by:

$$L_{t} = L_{t-\Delta t} \exp \left( f(\delta + t) - f(\delta + t - \Delta t) + \upsilon_{t} + \varepsilon_{t} - \varepsilon_{t-\Delta t} \right),$$
(11)

$$L_t = \tilde{L}_t \exp\left(\sum_{1}^t (\Pi_s + G_s)\right),\tag{12}$$

for  $t \in (0, T)$ . Here,  $L_0 = \widetilde{L}_0$  w.p. 1,  $\upsilon_t \sim \text{i.i.d. } N(0, \sigma_{\upsilon}^2 \Delta t)$  and  $\varepsilon_t \sim \text{i.i.d. } N(0, \sigma_{\varepsilon}^2)$ .

#### 4.5. Prices and yield curves on the scenario tree

The financial model set out above is used to generate the hybrid scenario structure as described in Sections 3.3 and 3.4. The ten state variables stored at each node are  $[R_n, \Pi_n, r_n, \pi_n, \beta_{1,n}, \beta_{2,n}, \beta_{3,n}]$ , employing the same notation as before except that we index by node *n*. Asset prices can be evaluated at each node of the scenario tree. For example, the asset price in Eq. (6) is transformed into the nodal form simply by replacing *t* with *n* and  $t - \Delta t$  with  $n^-$  as follows:

$$S_{i,n} = S_{i,n} \cdot \exp(R_{i,n}), \text{ for } n \in \{\mathcal{N}_t, t \in (0, T]\} \text{ and } i \in \mathcal{F},$$

with  $S_{i,n_0} = 1$ . Pricing formulas for other assets are transformed similarly.

It is useful to consider the yield curves generated by the scenario tree. Percentiles of the nominal and real spot rates for different maturities at retirement date *T*, as calculated from the generated scenarios, are shown in Table 5. It is clear that the generated nominal yield curve is more volatile at the short end, than at the long end, consistent with empirical evidence. This is also evident from the right hand panel of Fig. 2, which shows eight different realizations of the nominal yield curve, with a greater spread at the short end than at the long end.

#### 5. Numerical results

In this section, we solve the investment and deferred annuitization problem for an individual who can invest in equities, cash, inflation-indexed and nominal bonds, and who can buy both inflation-protected and nominal annuities. We investigate particularly the welfare enhancement potentially conferred by inflationprotected deferred annuities in the presence of labour income risk.

Table 5	5
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Percentiles of the nominal and real spot rates at retirement, for different maturities.

	Percentile	5y	10y	15y	20y	25y	30y
Nominal	0.05	-0.0563	-0.0406	-0.0302	-0.0233	-0.0182	-0.0148
Nommai	0.95	0.1212	0.1160	0.1109	0.1065	0.1030	0.1002
Real	0.05 0.50 0.95	-0.0550 0.0003 0.0557	-0.0473 0.0042 0.0556	-0.0458 0.0052 0.0560	-0.0460 0.0054 0.0567	-0.0464 0.0054 0.0573	-0.0468 0.0054 0.0576

#### 5.1. Numerical example

Consider a 40-year-old individual ( $\delta = 40$ ) who intends to retire at age 65 (T = 25).<sup>5</sup> His aim is to maximize and secure his retirement benefits *in real terms* and to set aside a portion of his portfolio as a bequest, also *in real terms*, if he dies before retirement. His personal retirement plan permits him to invest in an equity fund, a nominal bond fund (maturity M = 20 years), an inflationlinked bond fund (maturity M = 20 years), a cash fund (maturity M = 5 years), nominal deferred annuities, and inflation-protected deferred annuities as described in Section 2.

In the base case, the VAR asset return model with Nelson–Siegel real and nominal yield curves is based on U.K. data from January 1985 to June 2017, as described in Section 4.2. Mean, standard deviations and cross-correlations of key variables from the model are shown in the online supplementary section, Table S-1. A U.K. mortality table based on 2000–2006 experience<sup>6</sup> is used to price the deferred annuity.

The individual can rebalance his portfolio and buy deferred annuities every 5 years ( $\Delta t = 5$ ) until retirement. There are therefore six stages (five periods) in the scenario tree part of the hybrid scenario structure. At time 0, the individual has wealth of  $w_0 = \pm 80,000$  in his retirement plan. His annual wage starts at  $\pm 40,000$ , so  $L_0 = \widetilde{L}_0 = 40,000$  in Eqs. (11) and (12), or equivalently  $L_{n_0} = \widetilde{L}_{n_0} = 40,000$  in Eqs. (S.5) and (S.6). He contributes ten per cent ( $\phi = 10\%$ ) of his nominal labour income to his retirement plan.

Average real labour income with age typically exhibits a concave shape unless deflation occurs. National statistics show that the average real wage tends to decline after age 40–50 (Office for National Statistics, 2015). We employ a deterministic quadratic function for f in the real labour income mode of Eq. (11), and use the parameter estimates of Blake et al. (2007)<sup>7</sup>.

## 5.2. Optimal annuitization and investment without risk to real labour income

First, we consider the case where there is no real labour income uncertainty: there are no stochastic shocks to real labour income,  $\sigma_{\upsilon} = \sigma_{\varepsilon} \equiv 0$  in Eq. (11).

Our numerical results show that the optimal strategy to secure retirement income leads to fairly early and continual purchase of deferred annuities over the working lifetime of the investor. This is illustrated in Fig. 3, which shows average and percentile values of annual retirement income accumulated in real terms from holdings of nominal deferred annuities (N-DAs) and inflation-protected deferred annuities (IP-DAs), at various ages. From age 65 onwards, the annual annuity payouts, in real terms, are shown and these are constant for the infation-protected annuity but declining for the nominal annuity. The four panels represents four different individuals with different risk aversion and time preferences. There is no bequest motive ( $\kappa = 0$ .).

It is clear from Fig. 3 that nominal annuities are purchased at earlier ages than inflation-protected annuities. This is because a younger investor implicitly holds larger human capital than an older one *ceteris paribus*, labour income is perfectly correlated with inflation here, so the nominal annuity is riskier than the inflationprotected annuity, and at younger ages more risk can be taken. We also observe, by comparing the panels on the right to the ones on the left in Fig. 3, that a more risk-averse investor buys less nominal annuities and more inflation-protected annuities, the latter being less risky than the former in real terms. Comparing the bottom panels to the top ones in Fig. 3, we see that the more impatient an investor, the more he purchases nominal annuities. This is sensible since cash flows from nominal annuities decline in real terms, satisfying the more impatient investor.

Average optimal asset allocations over period till retirement are presented in Fig. 4. The bar graphs in the first and third rows show average asset allocations including deferred annuities, and those in the second and fourth rows concentrate on financial assets only, excluding deferred annuities. At retirement (age 65), only annuities are held. Comparing the right-hand panels to the left-hand panels, we find that equity allocation is lower, bond allocation higher, and annuity allocation marginally higher, the greater the investor's riskaversion is.

A key result from Fig. 4 is the dominance of deferred annuity holdings over financial asset holdings. This is a function of the historical data, 1985–2017, that we used to parameterize our financial market model and price annuities. Long-term rates fell during this period, meaning that annuities were cheap at the start, and our model rightly favours annuities over financial assets, on average.

We also find that nominal deferred annuities dominate their inflation-protected cousins, and that nominal bonds dominate inflation-linked bonds. This initially surprising result follows from the fact that long-term real yields were on average negative over this historical period (the mean of  $\tilde{\beta}_{1,t}$  in online supplementary section Table S-1 is -0.12%) so inflation-indexed securities and products were expensive. Over this historical period, the longterm nominal yield is highly correlated with the long-term real yield (the correlation of  $\beta_{1,t}$  with  $\tilde{\beta}_{1,t}$  in online supplementary Table S-1 is 0.7319.). Investing in the long-term nominal bond fund thus helps to hedge price changes in not only nominal but also inflation-protected annuities to secure real retirement income. The average long-term nominal rate is 3.92% (the mean of  $\beta_{1,t}$  in online supplementary section Table S-1), which is higher than the average annual equity return (the average of monthly return  $r_t$  in online supplementary section Table S-1 is 0.22% accumulating to 2.63% p.a.), explaining the preference of nominal bonds over equities.

<sup>&</sup>lt;sup>5</sup> In the numerical results that follow, we assume that the individual has a CRRA (power) utility, as defined in section 2, with representative relative risk aversion coefficient  $\gamma$  of 3 or 5, and time impatience  $\rho$  of 0.02 or 0.04. This is in common with other long-term portfolio and annuity studies, e.g. Huang et al. (2017), Horneff et al. (2010), Koijen et al. (2011) and Viceira (2001).

<sup>&</sup>lt;sup>6</sup> Institute and Faculty of Actuaries, S1PML/S1PFL All pensioners (excluding dependants), male/female lives www.actuaries.org.uk/research-and-resources/ documents/s1pml-all-pensioners-excluding-dependants-male-lives (www. actuaries.org.uk).

<sup>&</sup>lt;sup>7</sup> The real labour income function scaled to 1 at the final age of 60 is  $w(y) = 0.5963 + 2.3708y - 1.9671y^2$ , where y = (x - 20)/(60 - 20), x = 20, 21,..., 60. The coefficients, estimated from the New Earnings Survey of the Office for National Statistics in 1998, are for a male all-occupation group. For details, refer to Blake et al. (2007).



**Fig. 3.** Average and percentiles of optimal secured retirement income ( $\times$ £1000 p.a.) through either inflation-protected deferred annuities (IP-DAs) or nominal deferred annuities (N-DAs) for different coefficients of risk aversion  $\gamma$  and time preference  $\rho$ , in the absence of labour income risk and bequest motive. All values are presented in real terms.

5.3. Optimal annuitization and investment with risk to real labour income

Next, we consider the case where there are risks to real labour income. We use the parameterization of Cocco et al. (2005), based on income data for a college-educated worker, for the volatility of permanent and temporary shocks to real labour income:  $\sigma_v^2 = 0.02$  and  $\sigma_{\varepsilon}^2 = 0.05$  in Eq. (11).

Fig. 5 compares secured retirement income without and with labour income risk. More retirement income is received from the inflation-protected annuities than from the nominal annuities when labour income risk is present. Fig. 6 compares the average allocations in the two cases. Allocations to inflation-protected annuities and to bonds are marginally higher in the presence of labour income risk than in its absence. These results are as anticipated: implicit wealth, including human capital, is riskier if the investor has risky real labour income than if he does not, so in the first case he should choose less risky investment and annuity strategies in terms of achieving real retirement income.

#### 5.4. Investigating the dynamic solution

The optimal investment and annuity strategies are, of course, dynamic. In this section, we investigate the dynamic strategy by considering the decisions made on specific scenario nodes. We demonstrate how the strategy is explained by realized asset prices on a node and expected asset returns on descendant nodes. The model parameters and assumptions are the same as in the case with labour income risk in Section 5.3.

Table 6 shows optimal investment and deferred annuity decisions on the root node at age 40 and on selected nodes at age 45, so that we can compare different scenarios. Case (a) in Table 6 shows that, at age 40, the optimal decision is to buy 12.55 units of the bond fund and 20.57 units of nominal deferred annu-

ity. The prices of the cash, bond, equity, and inflation-linked bond fund units are £1000. The prices of the nominal and inflation-protected deferred annuities are £4251 and £15,105 for £1000 p.a. of nominal and real retirement income, respectively.

Case (*b*) in Table 6 displays the scenario at age 45 with the lowest prices of nominal bond and nominal deferred annuity among all the scenarios at age 45. The lowest price of the nominal deferred annuity leads the risk-averse investor optimally to buy 9.379 units of this annuity. He also buys 31.653 units of the nominal bond fund because bond returns will mean-revert and this bond fund has an expected risk premium over cash of 16.3% over the next 5 years till age 50. All other financial assets are expected to have negative risk premiums, and are not purchased.

Case (c) in Table 6 presents the scenario at age 45 where the most units of nominal deferred annuity are bought. The price of this annuity has fallen (£3740 compared to £4251 in case (a)) whilst the price of the nominal bond fund fund has increased (£1172, up from £1000 in case (a)). The investor therefore sells about two thirds of his original holding of the bond fund at age 40 to buy nominal deferred annuities thereby securing £10,796 p.a. of nominal retirement income.

Case (d) refers to the scenario where the most units of the cash fund and inflation-protected deferred annuities are purchased. Cash is expected to return 37.9% over the next 5 years whereas the other assets are expected to return less, hence the large investment in the cash fund. The inflation-protected deferred annuity is cheap, having more than halved in price from case (a) from £15,105 to £7108, encouraging the investor to secure real retirement income of £2659 p.a. Similarly, the largest purchase of nominal bond in case (e) may be explained by its relatively cheap price and by its high expected return, with an expected premium of 11.1% over cash over the next 5 years. On the other hand, equity is the only asset whose price falls from case (a) to the scenario in case (f), and it also has the highest expected risk premium among the financial





**Fig. 4.** Optimal investment and deferred annuity proportions on average at different ages till retirement for different coefficients of risk aversion  $\gamma$  and time preference  $\rho$ , in the absence of labour income risk and bequest motive. The second and fourth rows show proportions of financial assets only. IL-Bond is inflation-linked bond, IP-DA is inflation-protected deferred annuity, N-DA is nominal deferred annuity.

assets. Unsurprisingly, case (f) is the scenario at age 45 with the largest equity purchase.

The dynamic optimal annuity and investment strategy is therefore explained by current asset prices and expected returns. The predictability of asset returns, in the VAR model fitted to data in Section 4, is exploited by the multi-stage stochastic programming

age

model, and the optimal strategy dynamically responds to market changes.

age

#### 5.5. Availability of deferred annuities

We wish to compare the situations where the availability of different types of annuities is restricted. We consider six possible



**Fig. 5.** Average and percentiles of optimal secured retirement income ( $\times$ £1000 p.a.) through either inflation-protected deferred annuities (IP-DAs) or nominal deferred annuities (N-DAs) in the absence of labour income risk (left panel) and with labour income risk (right panel). Risk aversion  $\gamma = 5$  and time preference  $\rho = 0.02$  are the same. All values are presented in real terms. There is no bequest motive.



**Fig. 6.** Optimal investment and deferred annuity proportions on average at different ages till retirement in the absence of labour income risk (left panel) and with labour income risk (right panel). Risk aversion  $\gamma = 5$  and time preference  $\rho = 0.02$  are the same. There is no bequest motive. IL-Bond is inflation-linked bond, IP-DA is inflation-protected deferred annuity, N-DA is nominal deferred annuity.

#### Table 6

Optimal investment and deferred annuity choices with labour income risk at ages 40 and 45 under different scenarios. Risk aversion and time preference are  $\gamma = 5$  and  $\rho = 0.02$ . No bequest motive.

Age	Node	Conditional expected 5-year returns [risk premium in brackets]			Net purchases of units <sup>a</sup> (prices in parentheses)						
		Cash	Bond	Equity	IL-Bond	Cash	Bond	Equity	IL-Bond	IP-DA <sup>b</sup>	N-DA <sup>b</sup>
40	(a) Root node	0.113	0.266 [0.153]	0.134 [0.021]	0.199 [0.086]	0.000 (1000)	12.555 (1000)	0.000 (1000)	0.000 (1000)	0.000 (15,105)	20.570 (4251)
45	(b) The lowest bond and N-DA prices	0.394	0.556	0.369 [-0.025]	0.351 [-0.043]	0.000 (1,113)	31.653 (596)	0.000 (2485)	0.000 (889)	0.000 (6703)	9.379 (1816)
	(c) The largest N-DA purchase	0.111	0.330	0.273	0.217	0.000 (1113)	-8.360 (1172)	0.000	0.000	0.000	10.796 (3740)
	(d) The largest cash and IP-DA purchases	0.379	0.285 [-0.094]	0.214	0.302	8.218 (1,113)	-12.555 (709)	0.000 (519)	0.000 (774)	2.659 (7108)	0.000 (2896)
	(e) The largest bond purchase (f) The largest equity purchase	0.129 -0.194	0.239 [0.111] -0.203 [-0.009]	0.125 [-0.004] -0.139 [0.055]	0.205 [0.076] -0.140 [0.054]	0.000 (1113) 0.000 (1,113)	34.029 (1184) -12.555 (2412)	0.000 (949) 38.409 (800)	0.000 (1149) 8.743 (1959)	0.000 (14,143) 0.000 (44,020)	0.000 (5262) 0.000 (16,742)

<sup>a</sup> Units in funds, or units of deferred annuities. A negative number means that there is a net sale.

<sup>b</sup> N-DA (IP-DA) = nominal (inflation-protected) deferred annuity. One unit of N-DA (IP-DA) is equivalent to £1000 p.a. of secured nominal (real) retirement income.

situations depending upon which feature is unavailable: nominal or inflation-protected payments, deferred or immediate payout. To compare these situations, we calculate a certainty equivalent (CE) value, which is the level of constant real retirement income which generates a utility equal to the maximized expected utility, in each of these situations. We use the numerical example in Section 5.3, i.e. with labour income risk. The results are tabulated in Table 7. The worst CE value is recorded when only nominal immediate annuities are available irrespective of the investor's risk and time preferences (last column of Table 7 labelled N/IA). Unsurprisingly, when there is utmost flexibility and all types of annuities are available, the highest CE value is recorded (third column of Table 7 labelled IP+N/DA+IA), but this is about 60% better in terms of equivalent real retirement income



**Fig. 7.** Optimal investment and deferred annuity allocations of overall wealth on average over the planning horizon and for different datasets. Constant relative risk aversion ( $\gamma$ ), time preference ( $\rho$ ), and bequest ( $\kappa$ ) coefficients are  $\gamma = 5$ ,  $\rho = 0.02$ , and  $\kappa = 0.0$  respectively. Volatilities of the permanent and temporary shocks are 0.02 and 0.05 respectively.

#### Table 7

Certainty equivalent values (£ p.a.) when the availability of different types of annuity is restricted, for different coefficients of risk aversion  $\gamma$  and time preference  $\rho$ .

Preferences		Availability									
γ	ρ	IP+N DA+IA	IP DA+IA	N DA+IA	IP+N IA	IP IA	N IA				
3	0.02	14396.83 (0.00%)	9316.08 (-35.29%)	13801.32 (-4.14%)	12567.78 (-12.70%)	8911.84 (-38.10%)	8424.12 (-41.49%)				
3	0.04	20551.62 (0.00%)	13025.18 (-36.62%)	20003.43 (-2.67%)	17751.67 (-13.62%)	12459.99 (-39.37%)	12206.63 (-40.60%)				
5	0.02	13227.98 (0.00%)	10180.07 (-23.04%)	11990.06 (-9.36%)	11250.87 (-14.95%)	9458.91 (-28.49%)	8250.79 (-37.63%)				
5	0.04	15976.37 (0.00%)	12037.21 (-24.66%)	14785.08 (-7.46%)	13391.44 (-16.18%)	11184.49 (-29.99%)	10167.68 (-36.36%)				

**IP+N/DA+IA**: both inflation-protected and nominal annuities are available, in both deferred and immediate versions; **IP/DA+IA**: only inflation-protected annuities are available, both deferred and immediate; **N/DA+IA**: only nominal annuities are available, both deferred and immediate; **IP+N/IA**: both nominal and inflation-protected annuities are available, but only immediate annuities; **IP+N/IA**: only inflation-protected immediate annuities are available, but only immediate annuities; **IP/IA**: only inflation-protected immediate annuities are available; **N/IA**: only nominal immediate annuities are available.

than the worst case. This suggests that individual investors' welfare is severely impaired in the current U.K. markets where the most common distribution strategies offered are nominal immediate annuities and income drawdown.

An interesting finding is that inflation-protected immediate annuities enhance welfare compared to their nominal cousins (compare the last two columns of Table 7) but that, when deferment is available, access to a nominal annuity is better than access to an inflation-protected annuity (compare the columns labelled IP/DA+IA and N/DA+IA). As we saw in Section 5.2, the average longterm real interest rate was slightly negative in the historical period (1985–2017) on which our model is calibrated, making inflationprotected deferred annuities very expensive.

These results show the importance of the availability of deferred annuities and also that, for long deferment periods, providing a nominal deferred annuity may be better than an expensive inflation-protected deferred annuity.

#### 5.6. Model validation and results over different historical periods

In order to validate our results, we generate several scenario structures to verify that our optimal solutions do not vary significantly both qualitatively and quantitatively. For any given historical period over which our model is parameterized, we only show the results on one scenario structure so that comparisons do not involve sampling error.<sup>8</sup> However, we also calibrate the scenario

<sup>&</sup>lt;sup>8</sup> We thank a reviewer for suggesting that a number of scenario trees can be generated and used for formal validation purposes. For robustness, results can also be averaged over several scenario trees.

structure over different historical periods in order to test that our central conclusions hold, and we describe this here.

In the earlier sections, our financial market model was parameterized on U.K. data in the period 1985–2017. Because of the negative average long-term real yield, and the steep upward sloping nominal yield curve, nominal bonds and nominal deferred annuities played a significant role in the average asset allocation. In this section, we calibrate our model on different historical sub-periods to examine if investment and annuity decisions change. Average allocations are shown in Fig. 7 for four different periods: the original Jan. 1985–Jun. 2017 period, Jan. 1985–Dec. 1997 (before the Asian financial crisis), Jan. 1997–Dec. 2007 (between the Asian and the global financial crisis), and Jan. 2008–Jun. 2017 (after the global financial crisis).

Average allocations are strikingly different over these different periods. First, this reminds us that our model produces *dynamic* asset allocation and annuitization decisions, but that only averages are shown in Fig. 7. Second, optimal portfolio decisions are sensitive to model calibration. Third, deferred annuities are bought early and in increasing amounts, on average, in all four periods in Fig. 7, and thus have a significant role to play in retirement planning, irrespective of the economic cycle.

We note that our model produces sensible results in each subperiod considered. For example, in the 2008–17 sub-period (bottom right panel of Fig. 7), negative long-term real yields and nearzero nominal yields make deferred annuities and bonds expensive and unappealing, and the model takes advantage of the rising stock market with a large and falling equity allocation. In the 1985–97 sub-period (top right panel of Fig. 7), high short rates to combat inflation make cash attractive in the earlier years. This is discussed further in online supplementary appendix Section S-4.

#### 6. Conclusion

We construct an optimal investment model with deferred annuities for an individual investor who is saving for retirement. The individual's labour income is uncertain and is subject to exogenous permanent and temporary shocks in addition to price inflation. Our results show that buying nominal and inflation-protected deferred annuities, continually and from an early age, is an optimal strategy when the objective is to maximize the expected utility of real retirement income and when the retirement income is secured by annuitization. The balance between inflation-protected and nominal deferred annuities depends on expectations of long-term real and nominal rates. In the presence of shocks to real labour income, inflation-protected deferred annuities are marginally preferred to nominal deferred annuities.

Optimal investment in financial assets (cash, nominal bond, inflation-linked bond and equity) also depends on financial market expectations, but portfolio riskiness typically declines with age, consistent with life-cycle models. If welfare is measured in terms of a certainty equivalent of real retirement income, we find that welfare falls by about 40% if deferred annuities are not available and only nominal immediate annuities are available at retirement. When we calibrate our model on different historical sub-periods, we find that our key result about the importance of early purchases of deferred annuities appears to be robust to different financial market expectations. The actual portfolio composition and annuity allocation vary, but the investment portfolio becomes less risky as retirement approaches.

Besides the results that we obtain in the context of retirement planning, we make some contributions in terms of modelling using multi-stage stochastic programming. We suggest a new hybrid scenario structure which combines a scenario tree in the pre-retirement phase with scenario fans in the post-retirement phase. We also implement an unconstrained non-linear program to generate shocks to real labour income which are independent of financial market scenarios. We also implement a scenario generation procedure that avoids arbitrage, using the methods of Klaassen (2002) and Geyer et al. (2014).

Future work will address some of the limitations of our model. Jurisdiction-specific taxes and transaction costs can be implemented in the model to enhance its practical usefulness. Retirement provision also takes place through employer-sponsored institutional vehicles, which deserve further investigation (Consiglio, Tumminello, & Zenios, 2015). Flexibility in the labour supply, i.e. hours worked, affects contributions to the retirement plan and is not considered here. The effects of housing and mortgage costs, and different mortality rates for impaired lives or different population groups, are also ignored. Power utility does not capture the elasticity of intertemporal substitution in consumption and an Epstein-Zin utility may be implemented. Other rational and behavioural factors such as habit formation and hyperbolic discounting can also be considered. Annuity products such as variable annuities can be added to the portfolio of annuities available. Life and health insurance can also affect optimal retirement planning. These extensions will be studied in subsequent work.

#### Supplementary material

Supplementary material associated with this article can be found, in the online version, at doi:10.1016/j.ejor.2021.03.052.

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